

Visual cortex activation predicts visual preference: Evidence from Britain and Egypt

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Abstract

The term “Perceptual goodness” refers to the strength, obviousness, or salience of a visual configuration. Recent work has found strong agreement between theoretical, neural, and behavioural measures of perceptual goodness across a wide range of different symmetrical visual patterns. We used these pattern types again to explore the relationship between perceptual goodness and aesthetic preference. A group of 50 U.K. participants rated the patterns on a 0 to 100 scale. Preference ratings positively correlated with four overlapping measures of perceptual goodness. We then replicated this finding in Egypt, suggesting that our results reflect universal aspects of human preference. The third experiment provided consistent results with a different stimulus set. We conclude that symmetry is an *aesthetic primitive* that is attractive because of the way it is processed by the visual system.

Keywords

Aesthetics; holographic model; perceptual goodness; sustained posterior negativity; symmetry

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Introduction

Perceptual goodness

Perceptual goodness is a term from the early Gestalt school, referring to perceptual strength, obviousness, or salience of a visual configuration (Koffka, 1935; Wertheimer, 1923; for a recent review, see Wagemans, in press). Reflectional symmetry is a *good gestalt*, where the structure is immediately apparent to human observers. Other symmetries, like repetition and rotation, are less salient (Mach, 1886). Different visual regularities are shown in Figure 1. Note how some of these are more obvious than others.

There is no definitive list of rules governing perceptual goodness, although some have argued that it is linked to simplicity and redundancy (Attneave, 1954; Hochberg & McAlister, 1953; Pomerantz & Kubovy, 1986). Following this theme, Van der Helm and Leeuwenberg (1996) proposed their *holographic weight of evidence model*, which quantifies the perceptual goodness of different regular dot patterns. Their key formula states that $W = E/N$, where E is evidence for regularity and N is the total amount of information. For reflectional symmetry with a single fold, E is the number of mid-point collinear dot pairs across the axis and N is the total number of dots.

Consequently, W is always 0.5; however, many extra symmetrical pairs are added. For repetition, E is the number of repeated blocks minus 1 and N is again the number of dots. W goes down as we increase the number of dots, whereas W goes up if we increase the number of repeated blocks. For Glass patterns (Glass, 1969), E is the number of dipole dot pairs minus 1 and N is number of dots. W rapidly approaches 0.5 (the same as 1-Fold reflection) when the number of dipoles increases. The scope and assumptions of the holographic model have been debated (Olivers, Chater, & Watson, 2004; van der Helm & Leeuwenberg, 2004). However, the W goodness metric

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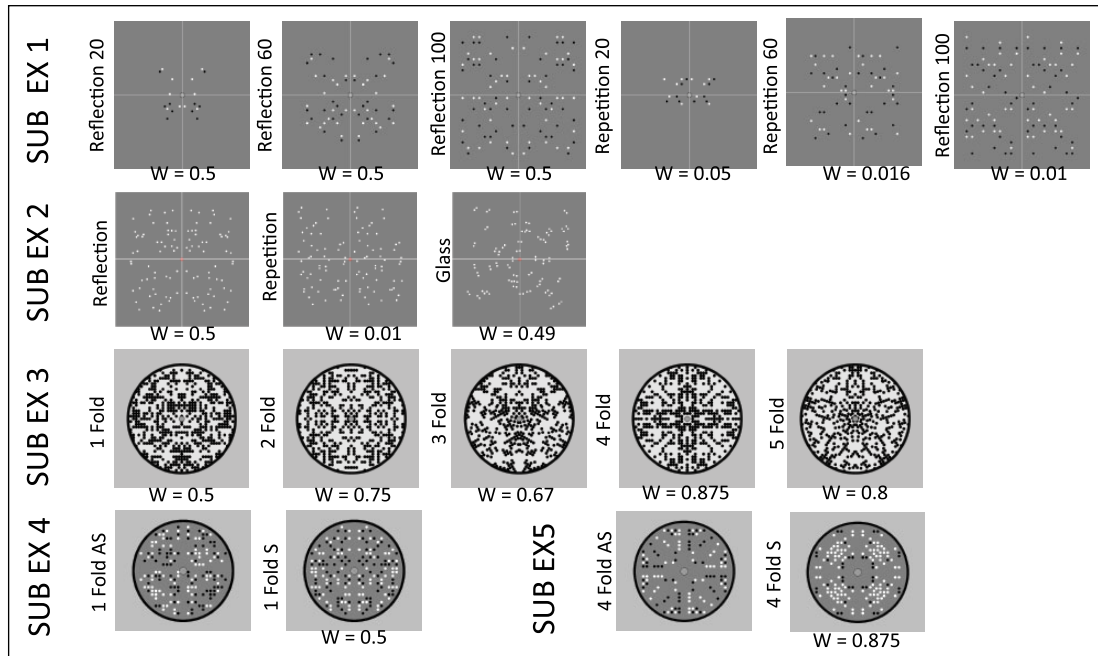


Figure 1. Examples of the regular patterns from our five sub-experiments. Participants completed all five experiments. Random patterns are not shown here, but available in Figure 1 of Supplemental Material 1. These are just examples. Different patterns were generated afresh on every trial, so no participant ever saw the exact same stimulus set. The same pattern types were used by Makin et al. (2016).

successfully predicts performance in regularity discrimination tasks (Nucci & Wagemans, 2007).

Makin et al. (2016) ran a series of five studies using regular patterns like those in Figure 1, where W ranged from 0.01 to 0.875 (and see Figure 1 of Supplemental Material 1 for more examples, including the matched random patterns). W explained substantial variance in grand-average response time in their forced-choice regular/random discrimination tasks ($r^2 = .88$), and W also explained considerable variance in grand-average error rate ($r^2 = .78$).

Visual symmetry generates an event-related potential (ERP) component called the *Sustained Posterior Negativity* (SPN; Figure 2 of Supplemental Material 1). Amplitude is more negative for symmetrical than random patterns, at posterior electrodes, from around 250 ms onwards (Bertamini & Makin, 2014; Makin, Wilton, Pecchinenda, & Bertamini, 2012; Norcia, Candy, Pettet, Vildavski, & Tyler, 2002). The SPN is generated by extrastriate visual regions, including the lateral occipital complex (LOC), where symmetry activations have also been detected with functional magnetic resonance imaging (fMRI) (Kohler, Clarke, Yakovleva, Liu, & Norcia, 2016; Sasaki, Vanduffel, Knutsen, Tyler, & Tootell, 2005). Makin et al. (2016) found that the W strongly predicted grand-average SPN amplitude ($r^2 = .86$).

We, thus have four separate estimates of perceptual goodness for the patterns in Figure 1. Two of these measures are behavioural (response time and error rate), one neural (SPN amplitude), and one theoretical (W load). All the four measures are potentially limited (response time

and error rate are noisy and distorted by ceiling effects, ERPs are noisy for different reasons, and W is noiseless but ignores some systematic influences on perceptual goodness). Nevertheless, these disparate measures of perceptual goodness are all strongly correlated (see scatterplots in Figure 3 of Supplemental Material 1). This gives a solid foundation for our work.

Scientific aesthetics

Early pioneers such as Birkhoff (1932) and Eysenck (1941) proposed formulas for relating objective stimulus features to aesthetic preference. This research programme continues (for review, see Palmer, Schloss, & Sammartino, 2013) but faces substantial challenges. For one thing, it is impossible to evoke strong emotions such as *intense fascination* (Kubovy, 2000) and *aesthetic rapture* (Markovic, 2012) with controlled stimuli on repeated trials. Furthermore, the human aesthetic faculty is tuned to gestalts, or wholes, and responses cannot be predicted by summing preference for parts (Holmes & Zanker, 2012). Some reliable effects have been discovered despite these limitations: Most people prefer symmetrical to random arrangements (Eisenman, 1967; Makin, Pecchinenda, & Bertamini, 2012), blue to brown (Palmer & Schloss, 2010), and smooth curvature to angularity (Bertamini, Palumbo, Gheorghes, & Galatsidas, 2016; Cotter, Silvia, Bertamini, Palumbo, & Vartanian, 2017).

This project examined the relationship between perceptual goodness and visual preference. People may like the

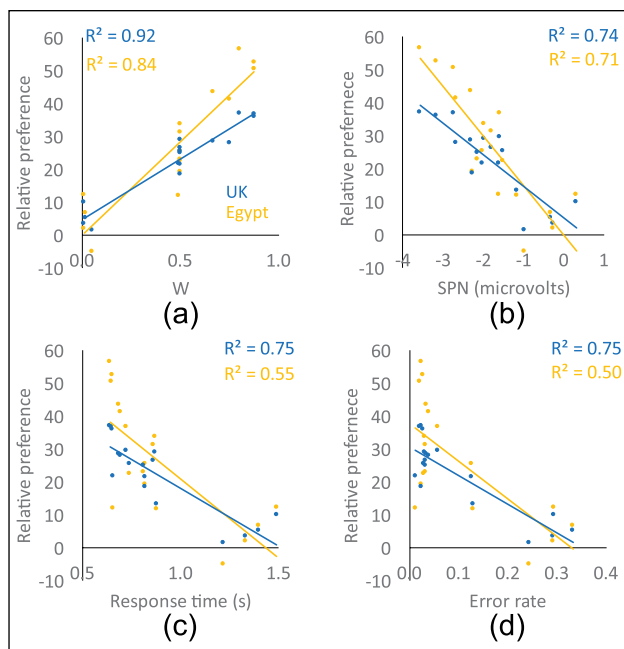


Figure 2. Relative preference correlates with four different measures of perceptual goodness. Note that a participant's *relative preference* is the difference between their mean ratings for regular patterns and mean ratings for random patterns. Positive values indicate a relative preference for regularity. Each data point represents the grand-average from one regular condition and sub-experiment. The variables on the X axis were obtained using the same types of patterns, but from different participants (Makin et al., 2016). (a) W load from the holographic model, (b) grand-average SPN amplitude, (c) grand-average response time, and (d) grand-average error rate. Note that an extremely negative SPN = a large neural response to symmetry and lower response time, and error rate = more efficient performance. Blue data are from the U.K. sample, and yellow data are from the Egyptian sample.

more obvious regularities, which have a high W load and generate a large SPN. This prediction has a long history—many scholars have claimed that beauty arises from the balance or combination of two fundamental factors. The first factor can be described as “*order, unity, or harmony*” and the second factor as “*complexity, multiplicity, or diversity*” (Boselie & Leeuwenberg, 1985; Eisenman, 1967; Roberts, 2007). Classic work proposed that beauty = order \times complexity (Eysenck, 1941) or that beauty = order/complexity (Birkhoff, 1932). In the holographic model, order is the number of holographic identities (E) and complexity is the number of elements (N). Birkhoff's theory, thus predicts that people should like high W patterns (because $W = E/N$).

More recent work also makes similar predictions. The *fluency-attribution model* states that people are sensitive to the efficiency of their own perceptual and cognitive operations and often like things that are processed fluently (Reber, 2012). The high W patterns are processed quickly and generate large neural response and should theoretically be preferred.

Artists often exaggerate distinctive features to an unrealistic level (Ramachandran & Hirstein, 1999). It is essential for animals to look good to potential mates and for flowers to look good to pollinating insects. This often means some form of phenotypic exaggeration, such as enlarged and brightly coloured tail feathers, petals, or other kinds of sexual dimorphism. Perhaps this applies more generally, and human aesthetic interest can be aroused by unusually high levels of visual excitation in the extrastriate symmetry network.

It is easy to find counter-claims. Resolution of perceptual ambiguity (Van de Cruys & Wagemans, 2011) or discovery of representational fit between visual and abstract layers (Palmer et al., 2013) can be more fundamental to artistic success than brute visual excitation. Second, some visual dimensions have an aesthetically optimal mid-point and unpleasant extremes (e.g., Berlyne, 1970; Redies, 2007; Spehar, Clifford, Newell, & Taylor, 2003). Finally, perfect symmetry may have a sterile rigidity about it, and many people may prefer imperfect symmetry disrupted by noise (Gartus & Leder, 2013; McManus, 2005).

Experiment 1 measured explicit preference for the pattern types in Figure 1. Whereas Makin et al. (2016) ran separate studies on different groups of participants, here conditions were presented as five *sub-experiments*, and *all participants completed all five sub-experiments*. Will people prefer the high W patterns (that are discriminated efficiently and produce large brain response) or will they like less obvious regularities (that are discriminated less efficiently and produce an intermediate brain response)? Different accounts in scientific aesthetics make diverging predictions, so we had no strong a priori hypothesis.

There were 100 participants in Experiment 1. The first group of 50 were English speakers recruited at the University of Liverpool, UK. The other 50 were Arab-speaking undergraduates at Menoufia University in Egypt. Cross-cultural comparisons are important and useful, but the comparison between Western and Arab culture is particularly interesting. Classic Islamic art has a greater emphasis on abstract geometry (Gonzalez, 2001), whereas classic Western art celebrates beauty in human faces and bodies. However, Soueif and Eysenck (1971) studied preferences for shape and found similar results in Britain and Egypt and this has been confirmed recently (Bode, Helmy, & Bertamini, 2017). Although analysis of just two populations will never be sufficient to support claims of universality, it does at least go some way to addressing the generalisability of our results.

Experiment 1

Method

In total, 50 participants were from the University of Liverpool (mean age = 28.2 years, min = 17, max = 58, 20 males, four left-handed) and 50 were from Egypt (mean age = 19.9, min = 19, max = 22, three males, 0 left-handed).

The Egyptian participants were younger than the U.K. sample on average, and there were a higher proportion of females. These confounds would be problematic if we were reporting large-cross cultural differences. However, results from United Kingdom and Egypt were similar.

Most participants were students or research staff. The study had local ethics committee approval and was conducted in accordance with the Declaration of Helsinki (revised 2008). In a previous unpublished experiment on the same topic (but without exact matching to the five studies of Makin et al., 2016), we found strong effects with 28 participants (minimum $r^2 = .69$). We were confident that 100 was an adequate sample size.

Each participant was involved in five short *sub-experiments*, matched to the five studies reported in Makin et al. (2016). Patterns types were identical to Makin et al. (2016) and like those shown in Figure 1. Individual patterns were generated afresh on each trial according to an algorithm programmed in Python using PsychoPy (Peirce, 2007) (see Open Science Framework for source code and raw data; <https://osf.io/3amna/>). No participant ever saw the same exact pattern twice, and patterns were never repeated across participants. On each trial, a pattern was shown for 3 s and then participants entered a preference judgement on a 0 to 100 Likert scale. The text above the scale read, "How much did you like that pattern?" and the extremes of the scale were labelled 0 = *not at all* and 100 = *extremely* (Figure 4 of Supplemental Materials 1).

In each sub-experiment, different examples of the regular patterns were presented six times, and the number of random patterns was balanced to match the total number of regular ones. In other words, half the trials were regular and half were random. This was consistent with the design of the ERP and behavioural experiments reported in Makin et al. (2016). Participants were not restrained in a chin rest, but pattern size was approximately $5^\circ \times 5^\circ$ visual angle (as in Makin et al., 2016). The order of the five experiments varied between participants, and the experiment lasted around 25 min in total.

Relative preference was computed in each sub-experiment and for each participant. This involved three steps: (a) mean preference for the six regular patterns in each condition was computed, (b) mean rating for all random patterns was computed, and (c) relative preference score was computed as the regular mean – random mean. Positive values indicate relative preference for regularity. Relative preference is a superior metric to absolute rating because it isolates aspects of preference exclusively related to W .

The correlation between grand-average relative preferences and W was assessed with Pearson's r . However, the holographic model does not stipulate how to calculate W for anti-symmetry like that used in Sub-experiments 4 and 5 (see van der Helm & Treder, 2009). We excluded these data points, so the W versus relative preference correlations were based on the other 16 conditions. The correlation between relative preference and three measures of perceptual

goodness from Makin et al. (2016) was similar. Here, we simply use the grand-average response times, error rate, and SPN amplitudes (although these were obtained from different groups of participants and conditions) as our predictor variables. These correlations were based on 18 data points because we have these measures for anti-symmetry. SPN amplitude was defined as regular-random at PO7/8 electrodes, from 300 to 400 ms post stimuli. This was chosen because the correlation between SPN amplitude and W peaked in this early window, although it was very strong throughout the traditional 300 to 1000 ms SPN interval.

Relative preference in the sub-experiments was analysed with repeated-measures analysis of variance (ANOVA). The Greenhouse–Geisser correction factor was applied whenever the assumption of sphericity was violated (Mauchly's W , $p < .05$).

Results

We first examined *relative preference* scores across all five sub-experiments. Note that relative preference is calculated for each participant and condition. It is the difference between the mean rating for a regular pattern and the mean rating from the matched random pattern. Positive values indicate preference for regularity. Figure 2 shows that there was a strong linear relationship between W and grand-average relative preferences ($r = .898$, $p < .001$). Relative preference was also correlated with SPN ($r = -.820$, $p < .001$), response time ($r = -.758$, $p < .001$), and error rate ($r = .739$, $p < .001$).

Figure 2 shows r^2 values for U.K. and Egyptian samples alongside each scatter-plot. This gives the proportion of variance in *grand-average* relative preferences explained by different measures of perceptual goodness. Our goodness metrics explained more variance in United Kingdom than in Egypt. However, we were mindful that analysis based on aggregated data leads to an overestimation of effect size (Brand & Bradley, 2012). We, thus, ran three additional analyses that take participant and trial level into account.

First, we ran an equivalent analysis on all individual participants. Each participant provided 18 relative preference scores. We then analysed this in the same way as the grand-averages. Figure 3a shows r^2 values from each participant organised cumulatively (with rare correlations with atypical sign coded as $r^2 = 0$). This graphical convention means that area below the line gives the proportion of variance explained by the predictor (and therefore the areas above the line are unexplained variance). Figure 3a shows this for the U.K. (blue) and Egyptian samples (yellow). Area under the line (variance explained) ranged between 0.33 and 0.54.

Second, we computed regression coefficients and p -values with linear mixed-effects analysis, which uses data from all trials and participants (LME4 library in R; Bates, Mächler, Bolker, & Walker, 2015). Relative preference was the dependent variable (DV). Country, Sub-Experiment, Participant, and Trial

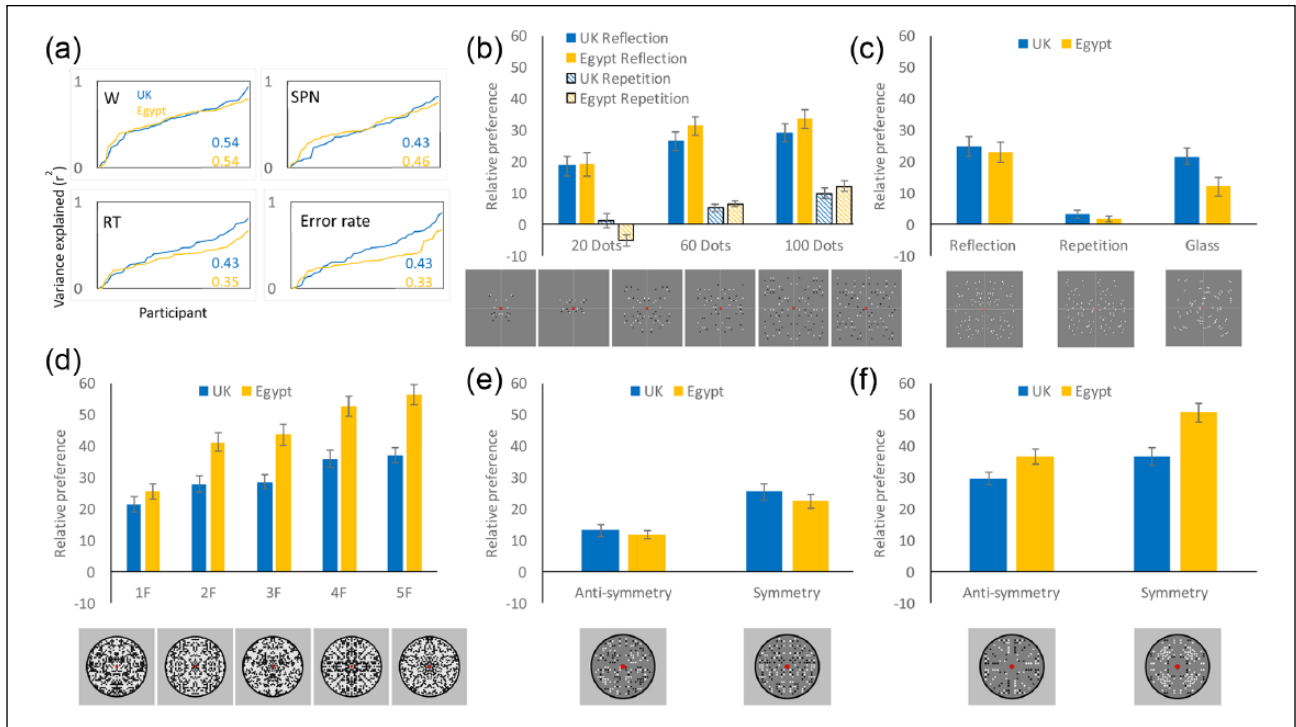


Figure 3. Results of Experiment 1. (a) Analysis of individual relative preference scores. Here, the proportion of variance (r^2) in relative preference was obtained from each participant and for each measure of perceptual goodness. The participants were then organised cumulatively. (b to f) Relative preferences of the five sub-experiments. Example patterns are shown in the insets below. Positive values indicate that the regular patterns of this type were liked more than random (maximum possible score = 100). Blue data are from the U.K. sample, and yellow data are from the Egyptian sample. Error bars = ± 1 SEM.

were included as random factors. This showed that an increase in W from 0 to 1 predicts a 44.45-point increase in relative preference (relative preference = $44.45W + 2.97$; $\chi^2(1) = 1,479.4$, $p < .001$). Strong effects were also found for SPN (relative preference = -10.01 SPN amplitude (μV) + 7.18 ; $\chi^2(1) = 1,226.1$, $p < .001$), response time (relative preference = -33.99 RT (s) + 53.60 ; $\chi^2(1) = 1,098.8$, $p < .001$), and error rate (relative preference = -78.44 error rate + 31.678 ; $\chi^2(1) = 1,373.5$, $p < .001$).

So far, the results fit a simple story of *Relative preference = Perceptual Goodness* and United Kingdom = Egypt. This was generally true when each of the five sub-experiments was analysed independently; however, there were a few systematic exceptions. Relative preference scores are shown in Figure 3b to f.

In Sub-experiment 1, there were main effects of regularity ($F(1, 98) = 167.690$, $p < .001$, $\eta_p^2 = .631$) and N-dots ($F(1.383, 135.522) = 47.077$, $p < .001$, $\eta_p^2 = .324$), but no Regularity \times N-dots interaction ($F(1.827, 179.063) = 1.596$, $p = .207$; Figure 3b). There would be an interaction here if relative preferences were perfectly predicted by W because N-dots has a unique effect on repetition according to the holographic model (van der Helm & Leeuwenberg, 1996). Furthermore, the main effect of N-dots is in the unpredicted direction - theoretically people should prefer repetition patterns with fewer dots. There

was also a weak N-dots \times Country interaction ($1.383, 135.522$) = 3.543 , $p = .048$, $\eta_p^2 = .035$) because the effect of N-dots was slightly weaker in the United Kingdom ($F(1.214, 59.501) = 19.376$, $p < .001$, $\eta_p^2 = .283$) than Egypt ($F(1.432, 70.178) = 28.235$, $p < .001$, $\eta_p^2 = .366$). There was no main effect of Country ($F(1, 98) < 1$).

The preferences in Sub-experiment 2 were closely related to W (Figure 3c). Participants liked reflection and Glass patterns almost equally and liked both far more than repetition ($F(2, 196) = 55.102$, $p < .001$, $\eta^2 = .360$). There was no main effect of Country ($F(1, 98) = 2.632$, $p = .108$) and no Regularity \times Country interaction ($F(2, 196) = 2.395$, $p = .094$).

According to the holographic model, there is a general increase in W with the number of axes of reflection, but 3- and 5-Fold reflections have slightly lower W loads than 2- and 4-Fold reflections (van der Helm, 2011). The dip at 3- and 5-Folds was not apparent in Sub-experiment 3, although there was not an increase either (Figure 2d). Relative preference generally increased with folds ($F(2.934, 287.526) = 110.582$, $p < .001$, $\eta^2 = .530$), and preferences were higher in Egypt than in the United Kingdom ($F(1, 98) = 14.508$, $p < .001$, $\eta_p^2 = .125$). There was a Folds \times Country interaction ($F(2.934, 287.526) = 11.486$, $p < .001$, $\eta_p^2 = .105$) because the effect of Folds was weaker in the United Kingdom ($F(2.709, 132.721) = 36.483$,

$p < .001$, $\eta_p^2 = .427$) than in Egypt ($F(2.861, 140.208) = 75.134$, $p < .001$, $\eta_p^2 = .605$).

In anti-symmetry patterns, black elements are paired with white and white elements are paired with black. Our Sub-experiments 4 and 5 measured preference for symmetry and anti-symmetry with 1- or 4-Folds (Figure 3e and f). The holographic model is silent about the goodness of anti-symmetry (because it is not considered to be a basic visual regularity; see van der Helm & Treder, 2009). However, we know that anti-symmetry is not discriminated as efficiently as symmetry under many conditions (e.g., Mancini, Sally, & Gurnsey, 2005) and sometime produces a slightly smaller SPN (Makin et al., 2016). Results for these sub-experiments are shown in Figure 2e and f. It is instructive to analyse these together. The participants liked 4-Fold patterns more than 1-Fold patterns ($F(1, 98) = 167.783$, $p < .001$, $\eta_p^2 = .631$) and symmetry more than anti-symmetry ($F(1, 98) = 102.316$, $p < .001$, $\eta_p^2 = .511$). There was no main effect of Country ($F(1, 98) = 2.681$, $p = .105$). However, there was a Folds \times Country interaction ($F(1, 98) = 16.748$, $p < .001$, $\eta_p^2 = .146$) and a three-way interaction between Folds, Regularity, and Country ($F(1, 98) = 8.891$, $p = .004$, $\eta_p^2 = .083$). For 1-Fold patterns, there was no Regularity \times Country interaction ($F(1, 98) < 1$) and no difference between United Kingdom and Egypt ($F(1, 98) < 1$). Conversely, for 4-Fold patterns, there was both a Regularity \times Country interaction ($F(1, 98) = 7.564$, $p = .007$, $\eta_p^2 = .072$) and higher preferences overall in Egypt ($F(1, 98) = 9.722$, $p = .002$, $\eta_p^2 = .090$).

Finally, we note that nearly every regularity was preferred to the matched random patterns presented in the sub-experiment. In other words, relative preference was always > 0 (one-sample t -tests, $p < .003$), with the sole exception of 20-dot repetition in Sub-experiment 1 ($t(99) = -1.196$, $p = .235$).

Discussion

Preference for different types of regularity increased with perceptual goodness (whether goodness is operationalised as W , SPN amplitude, response time, or error rate). Participants thus liked the most obvious regularities, which produced the largest SPN. This supports the idea that beauty = order/complexity (Birkhoff, 1932). Results are also consistent with the fluency account of aesthetics (Reber, 2012) and with the idea that people are attracted to visual exaggerations (Ramachandran & Hirstein, 1999). Our results suggest that symmetry is an *aesthetic primitive*: a visual property that is inherently interesting because of the way it is processed by the visual system (Latto, Brain, & Kelly, 2000). Although there were a few minor differences, results from the U.K. sample were closely replicated in Egypt.

An excessively bold conclusion here would be “Attraction to high W symmetry is a universal law of aesthetics, innately hardwired into visual and emotional brain

areas.” Of course, it would be naïve to make such a strong claim based on the available evidence. We therefore propose a nuanced version. Humans are hardwired so that some perceptual abilities always emerge whenever infants grow up in a typical visual environment. Through some combination of innate preparedness and exposure, the adult extrastriate cortex becomes better at discriminating some visual regularities than others (the holographic model provides a good estimate of adult sensitivity to different regularities). Preference ratings are linked to this regularity-sensitivity when regularity is the most prominent dimension in the data set. In the United Kingdom and Egypt, most students prefer the more salient regularities. This may generalise to most humans. More detailed proposals about evolution and development of symmetry perception and preference are included in van der Helm (2011).

Although the U.K. and Egyptian data sets were mostly similar, Egyptian participants tended to give higher ratings to the multiple symmetries. This might be because abstract geometry and multiple symmetries are celebrated in Islamic art (Gonzalez, 2001). Such effects could be superimposed on other universal laws of aesthetics (see Carbon, 2010 for analysis of the dynamics of fashion). However, there are other possibilities here as well. Perhaps everyone liked the multiple symmetries, but the U.K. sample was less inclined to use the high end of the response scale? Such limitations of the Likert scale procedure make us cautious about over-interpreting between-subject effects.

Precise predictions of the holographic model could be tested in future work. For instance, preference should increase with the number of repeated blocks in a repetition pattern and when the elements in a repetition can be grouped into salient blobs (Csathó, van der Vloed, & van der Helm, 2003).

Furthermore, the holographic model (van der Helm & Leeuwenberg, 1996) can be contrasted with the alternative *transformational model* (Garner, 1974). Among other things, these two models make different claims about the goodness of the N -Fold reflections used in Sub-experiment 3. Van der Helm (2011) quantified the predictions of the transformational model (which we denote T) as $1/(1-2F)$, where F is the number of folds. This means that T increases monotonically with F and asymptotes at high F . In contrast, the relationship between W and F is similar but *non-monotonic*, with a “goodness dip” at 3- and 5-Folds (Figure 1). Makin et al. (2016) tested whether W (holographic model) or T (transformational model) explained more variance in their DVs. T explained more variance in RT and error rate (although this can be explained by ceiling effects when $F > 2$). Conversely, W explained more variance in early SPN amplitude, and the predicted dip at 3- and 5-Folds was apparent. Previously, Wenderoth and Welsh (1998) found that discrimination of 3-Fold symmetry was often very like 2-Fold symmetry discrimination (and sometimes worse). Our Sub-experiment 3 found that relative preference for 2-Fold was approximately equal to 3-Fold and relative

preference for 4-Fold was approximately equal to 5-Fold. These results lie roughly between the predictions W and T (rather like the results of Wenderoth & Welsh, 1998).

We should also consider the generality of the Sub-experiment 3 results. All patterns had one vertical axis and 2- and 4-Fold reflections had additional horizontal axes, whereas 3-, 4-, and 5-Fold reflections had additional diagonal axes. The W scores for each of the N-Fold reflections would remain the same if they were globally rotated and all vertical and horizontal axes were eliminated. However, vertical and horizontal orientations might have a special status for the visual system (e.g., Wenderoth, 1994), so global rotation could potentially modulate preference independently of W . Indeed, it is certainly true that not all reliable goodness differences are captured by W differences. These anomalies can sometimes be plausibly explained as secondary influences on early visual processing (e.g., Csathó et al., 2003). This is also a topic for future work.

Experiment 2

So far, it seems that preferences for abstract symmetrical patterns are linearly related to perceptual goodness, as defined by the holographic model. However, our first

experiment did not test this directly. Our low and high W patterns were built from different elements and tested in different sub-experiments (Figure 1). In Experiment 2, we measured preference for symmetry+noise patterns which spanned the full W range. Example patterns are shown in Figure 4a. The proportion of symmetry varied between 0% and 100% in 20% increments. (We called this proportion “Psymm.”) There is a linear relationship between W and Psymm. These kinds of patterns were used in an ERP study by Palumbo, Bertamini, and Makin (2015), who found that SPN amplitude also scaled with Psymm (Figure 5 of Supplemental Material 1), whereas previous work has shown similar scaling of the extrastriate BOLD response and behavioural performance (Barlow & Reeves, 1979; Sasaki et al., 2005; also see van der Helm, 2010 for theoretical analysis). We, thus, expected that relative preference would increase with Psymm in Experiment 2.

Method

In total, 24 participants from the University of Liverpool were involved in Experiment 2 (aged 18–35 years, mean age=21.6 years, nine males, two left-handed). The preference rating protocol was the same as that

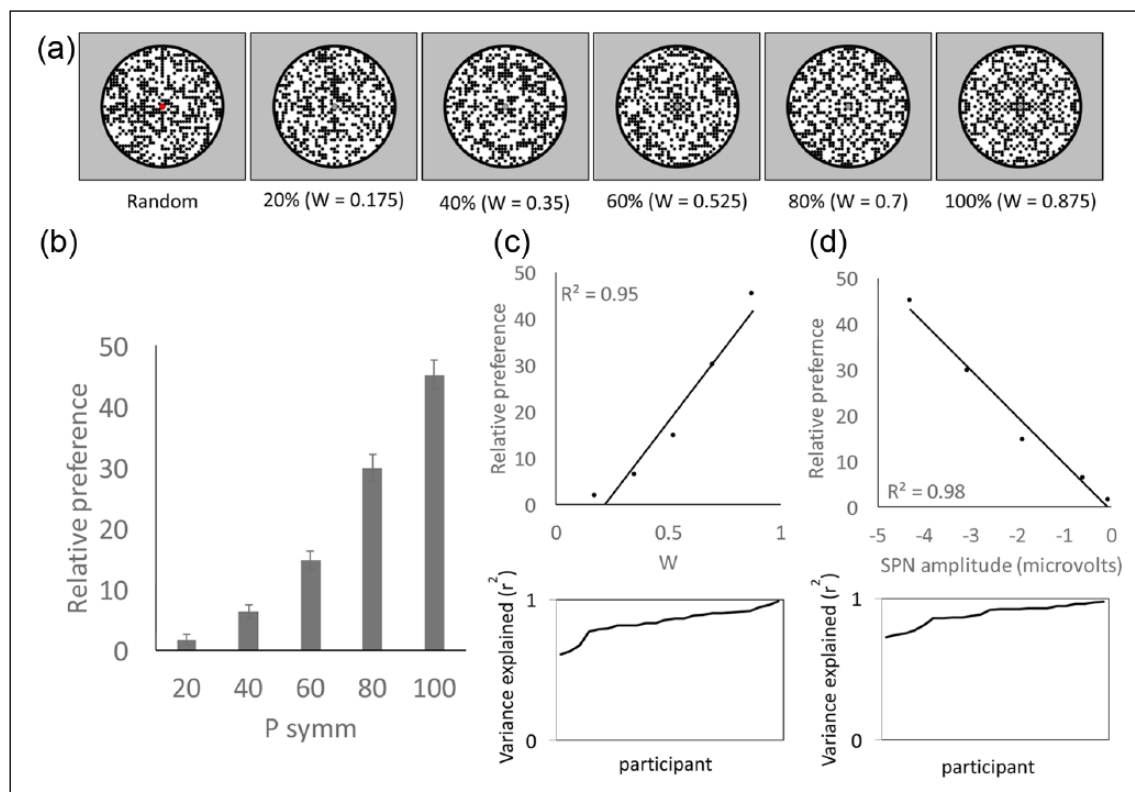


Figure 4. Stimuli and results of Experiment 2. (a) Example patterns from Experiment 2. These patterns are just examples. Novel examples were generated by the programme on every trial. Psymm and W loads are indicated below. (b) Relative preference as a function of Psymm (error bars = ± 1 SEM). (c) Correlation between W and relative preference, and r^2 values from Individual participants, organised cumulatively and (d) correlation between SPN amplitude and relative preference, and r^2 from Individual participants, organised cumulatively.

used in Experiment 1. There were 10 repeats of each Psymm level, and 50 random exemplars, giving 100 trials in total. The stimulus generation algorithm for these patterns is described by Palumbo et al. (2015) and considerations about W calculation are in Supplemental Material 1 of Makin et al. (2016). To illustrate, consider that an 80% symmetry is one where 80% of the dots are positioned symmetrically on a grid and the other 20% are positioned randomly (but might form accidental pairs across one or more of the axis). W was estimated by multiplying W for a 4-Fold reflection (0.875) by Psymm (e.g., $0.875 \times 0.8 = 0.7$) and ignoring accidental pairing.

Results and discussion

As expected, relative preference increased with Psymm ($F(2.039, 46.893) = 142.994, p < .001, \eta^2 = .861$), and the difference between each increment was significant ($p < .002$; Figure 4b). The five grand-average relative preference scores were almost perfectly correlated with W ($r = .98, p = .004$) and grand-average SPN amplitude ($r = .99, p = .001$). At an individual participant level, W explained 84% of variance in relative preference (Figure 4c) and SPN amplitude explained 88% of variance in relative preference (Figure 4d).

The results of Experiment 2 were thus very simple. Participants liked perfect symmetry, and preference increased linearly with Psymm. These results mirror other recent work by Gartus and Leder (2013), who also found that most participants preferred perfect symmetries to broken symmetries with slight imperfections. Supplemental Materials 2 provides further analysis of similarities and differences between preference and goodness metrics.

General discussion

Makin et al. (2016) found strong agreement between four different measures of perceptual goodness (W , SPN amplitude, response time, and error rate). In Experiment 1, we confirmed that participants gave higher preference ratings to the higher W patterns which produce larger SPNs. These effects were replicated in the United Kingdom and Egypt. Experiment 2 found that participants liked perfect regularity, which produces the largest SPN (and not regularity with a degree of noise).

We found that the relationship between goodness and preference is linear. There was no evidence that people liked mid-level goodness, where the structure was not too obvious. At least for the patterns used here, we can confidently assert that preference is a straight function of perceptual goodness. Preference for high W patterns is consistent with an aesthetic formula which states that beauty = order/complexity (Birkhoff, 1932). Given these results, we suggest that symmetry is an aesthetic primitive, that is, a feature

whose appeal derives directly from the way it is processed by the visual system (Latto et al., 2000).

Although we aimed to examine divergent accounts, we are mindful that our experiments cannot definitively falsify them. In scientific aesthetics, we must always be cautious about generalising claims far beyond the tested stimuli. Future experiments using other sets (e.g., real objects, faces, or art) might discover a different relationship between symmetry and preference. There may indeed be cases where perfect symmetry looks sterile and rigid (McManus, 2005) and other cases where it is preferred (Gartus & Leder, 2013). Indeed, there is some evidence that symmetry preferences are not uniform across categories (Little, 2014). For instance, painting or film where every visual detail was arranged symmetrically would look obviously wrong, even though compositional balance is often desirable (Arnheim, 1974). This kind of “gestalt nightmare” was analysed at length in Makin (2017).

Before moving on, we must consider one alternative explanation for our results. We assume that preference reports were strongly influenced by the perceptual goodness (perceptual goodness > preference > report). However, participants probably recognised that their own preferences were determined by perceptual goodness. They might have then taken a cognitive short-cut and reported perceptual goodness directly, rather than dwelling on how much they liked the patterns (perceptual goodness > report). Did our participants bypass their aesthetic faculties altogether and report perceptual goodness directly (perhaps mentally relabelling the response scale)? We cannot completely rule this out. However, previous studies have found a strong relationship between the salience of symmetry and implicit preferences (Bertamini, Makin, & Rampone, 2013; Makin, Pecchinenda et al., 2012), so it is unlikely that pure preference ratings (without the cognitive short-cut) would be completely different.

Many animals, including insects, fish, and birds, use phenotypic symmetry in mate selection and food choice, and humans often use symmetry to judge sexual attractiveness (Grammer, Fink, Møller, & Thornhill, 2003). However, we suggest that preference for abstract symmetry is NOT merely an overgeneralisation of innate mate selection strategies. Consider that multiple-axes symmetry is not face-like or body-like, but our participants liked it more than single-axis symmetry. Furthermore, Glass patterns are not at all face-like, but these were liked nearly as much as reflection. In summary, W is a far better predictor of preference than biological relevance. We propose that symmetry detection has a broad perceptual utility, and the appeal of symmetry is directly related to the strength of the symmetry signal in the extrastriate symmetry network.

Does this mean that other aesthetic accounts that emphasise ambiguity and resolution (Van de Cruys & Wagemans, 2011), representational fit (Palmer et al., 2013), and imperfection (McManus, 2005) are wrong?

We do not go that far: Instead, the results force us to think about the scope of different ideas in scientific aesthetics. Perhaps preference is directly linked to perceptual goodness when the stimuli are tightly controlled and when patterns are presented in a quasi-psychophysical lab experiment. Other accounts may describe the psychodynamics of aesthetic experience in the real world, where stimuli are multi-dimensional and aesthetic evaluation is optional and unconstrained. Even though scientific aesthetics is an old enterprise, it remains at an early stage of development. This kind of distinction is vital if we are to apply our theoretical insights correctly.


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Supplemental material

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